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ABSTRACT

The association between fluoride and risk for osteosarcoma is controversial. The purpose of this study was to determine if bone fluoride levels are higher in individuals with osteosarcoma. Incident cases of osteosarcoma (N = 137) and tumor controls (N = 51) were identified by orthopedic physicians, and segments of tumor-adjacent bone and iliac crest bone were analyzed for fluoride content. Logistic regression adjusted for age and sex and potential confounders of osteosarcoma was used to estimate odds ratios (OR) and 95% confidence intervals (CI). There was no significant difference in bone fluoride levels between cases and controls. The OR adjusted for age, gender, and a history of broken bones was 1.33 (95% CI: 0.56-3.15). No significant association between bone fluoride levels and osteosarcoma risk was detected in our case-control study, based on controls with other tumor diagnoses.

KEY WORDS: fluoride, osteosarcoma, case-control study, bone, oncology, epidemiology.

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An Assessment of Bone Fluoride and Osteosarcoma

INTRODUCTION

Osteosarcoma, a rare, painful, primary malignant bone tumor, is more prevalent in males (Homa *et al.*, 1991), in the long bones (Patel and Benjamin, 2005), and in individuals < 20 yrs old (Gurney *et al.*, 1999).

Chemicals and genetic factors have been suggested as risk factors of osteosarcoma (Miller *et al.*, 1996), while ionizing radiation is the only documented environmental risk factor for bone cancer (Steiner, 1965; Tucker *et al.*, 1987). A National Toxicology Program (NTP) study concluded that there was “equivocal evidence” of carcinogenic activity of sodium fluoride in male rats that were given extremely high doses (100 ppm and 175 ppm) for 2 yrs (NTP, 1990). Other animal studies have not provided evidence of an association between fluoride and osteosarcoma (Maurer *et al.*, 1990; NTP, 1992).

Numerous descriptive studies, with self-reported or ecological level data used to determine fluoride exposure from drinking water, failed to demonstrate an association (Hoover *et al.*, 1976; Doll and Kinlen, 1977, 1978; Newbrun, 1977; Hrudey *et al.*, 1990; Mahoney *et al.*, 1991; Freni and Gaylor, 1992). Similarly, case-control studies have not found any significant association between osteosarcoma risk and fluoridated drinking water (McGuire *et al.*, 1991; Moss *et al.*, 1995) or total lifetime fluoride (Gelberg *et al.*, 1995). One exploratory analysis reported an increased risk among a subset of males exposed to fluoride in drinking water during childhood (Bassin *et al.*, 2006).

Fluoride has an affinity for calcified tissues, with 99% of fluoride in the body contained within the skeleton. Thus, bone fluoride levels can serve as a biomarker for chronic fluoride exposure, providing a more objective measure of fluoride exposure. The purpose of this study was to evaluate

whether fluoride levels in bone are associated with the occurrence of osteosarcoma.

METHODS

Study Population

Patients were identified by physicians in the orthopedic departments from 9 hospitals across the US between 1993 and 2000. The study sample included incident cases of primary osteosarcoma, including osteoblastic, parosteal, and periosteal subtypes, and two control groups: tumor controls, patients with newly diagnosed malignant bone tumors; and orthopedic controls, surgical patients with benign tumors or non-neoplastic conditions. Since tumor controls were the only ones with available bone specimens for assay, they comprised the control series for this report. The study was approved by the Institutional Review Boards of the respective hospitals, Harvard Medical School, and the Medical College of Georgia.

All eligible patients who consented to participate were interviewed in person during hospitalization, pre-admission, or post-admission. Medical information was requested for all living patients born in the US. Patients who completed at least 80% of the questionnaire were considered to be enrolled in the study. Although the study protocol called for matching of cases and controls based on gender, age (± 5 yrs), and distance from their medical center, this approach was abandoned early in the study, since it proved to be a barrier to recruiting controls. Thus, all available tumor patients were recruited, and the statistical analysis was adjusted for age and gender.

Exposure and Outcome Assessment

Cancer diagnoses were confirmed by pathology reports. Specimens of both the tumor and normal bone adjacent to the margins of tumor tissue, herein referred to as tumor-adjacent bone, were collected from cases and tumor controls during surgery. Given that bone at the tumor site was destroyed as a result of the tumor, tumor-adjacent bone was analyzed for fluoride content. In some centers, a segment of bone from the iliac crest was also requested for cases, to assess the correlation between fluoride in iliac crest bone and in tumor-adjacent bone.

Methods used to measure fluoride concentration in the bone specimens have been described in detail elsewhere (Medina *et al.*, 2006). A 4- to 6-mg portion of bone was ashed, pulverized, and analyzed for fluoride concentrations (ash weight, mg F/kg, or ppm) according to a method developed by Taves (1968) and modified by Whitford (1996). With blinding to the case or control status of the bone specimens, each specimen was analyzed in duplicate; if measurements differed by more than 10%, another specimen was analyzed. Deer bone specimens with known fluoride concentrations were included in each batch of specimens for quality control (Medina *et al.*, 2006) and confirmed the validity of the bone fluoride assay procedure.

Statistical Analysis

We used Chi-square and Wilcoxon rank-sum tests to evaluate differences in patient characteristics and median fluoride

concentrations between tumor-adjacent bone and iliac crest bone. We also evaluated all specimens of tumor-adjacent bone and iliac crest bone among the cases, taking into account the within-person correlation for those patients who had both types of bone specimens. We used Spearman's correlation to assess the correlation between fluoride iliac crest bone and fluoride in tumor-adjacent bone.

In the subset of matched cases and controls, we used Wilcoxon's signed-rank test to evaluate if there were a difference in the median fluoride concentration in tumor-adjacent bone. Since this subset represented less than 25% of cases that provided bone, an unmatched analysis comparing median fluoride concentration in tumor-adjacent bone among all cases with control bone was conducted by a Wilcoxon rank-sum test.

We used both conditional and unconditional logistic regression to estimate the age- and sex-adjusted odds ratios (OR) and 95% confidence intervals (CI) to account for the initial matching. Fluoride measurements were transformed to a natural logarithmic scale to improve normality (Pagano and Gauvreau, 2000). Age- and sex-adjusted analysis was carried out for variables that were considered to be potential confounders of osteosarcoma: race/ethnicity; patient's, mother's, and father's education; combined household income; whether the patient ever lived in an urban area; and patient's past medical history (history of broken bones, other bone diseases, other cancers, receiving radiation for diagnosis or treatment prior to the present diagnosis); and variables with p value of ≤ 0.25 in the demographic-adjusted analysis were considered as potential confounders. The missing indicator method was used for patients missing information on household income and parents' level of education (Greenland and Finkle, 1995). Both manual and automated stepwise selection approaches were used to determine potential confounders to be included in the final risk-adjusted model. An exploratory analysis was also conducted among patients < 45 yrs old and < 20 yrs old. However, this study did not have sufficient power for a subgroup analysis among patients < 20 yrs old. Statistical analysis was carried out in SAS Version 9.1 (SAS Institute, Cary, NC, USA).

RESULTS

In total, 314 patients were eligible for enrollment (200 cases; 114 controls), and 296 patients (94%) completed the questionnaire (188 cases; 108 controls). Of these, 194 patients (142 cases; 52 controls) provided either tumor-adjacent or iliac crest bone for assay of fluoride content in bone. Eighteen patients were deceased, did not complete the questionnaire, or were otherwise lost to follow-up (12 cases; six controls). In total, 257 bone specimens were analyzed for fluoride content (200 tumor-adjacent bone; 57 iliac crest bone). Bone from six patients (five cases; one control) had fluoride levels below 100 mg F/kg, and thus were considered to be tissue other than bone (G. Whitford, personal communication).

Among patients who provided bone, there were no differences between cases and controls in enrollment site, race/ethnicity, patient's and mother's education level, combined household income, and whether they ever lived in an urban area. The median age of controls was higher than that of cases ($p < 0.001$);

gender-specific age differences were also significant, with controls being older, on average, than cases for both males ($p < 0.001$) and females ($p = 0.02$) (Table 1). There was a greater proportion of male cases than female cases ($p = 0.03$), and fathers of cases were significantly more likely to have higher education levels than those of controls ($p = 0.02$) (Table 1). Comparisons between all subjects who provided bone specimens versus those who did not are included in the Appendix Table 1.

Among the 53 cases who provided both tumor-adjacent and iliac crest bone specimens, there was a significantly higher fluoride concentration in iliac crest bone than in tumor-adjacent bone (median = 697 vs. 558 mg F/kg bone ash, $p < 0.001$) (Fig., a). However, when all specimens of tumor-adjacent bone ($N = 137$) and iliac crest bone ($N = 54$) from cases were included, the median fluoride concentrations in iliac crest bone were not significantly higher than those in tumor-adjacent bone (median = 695 vs. 611 mg F/kg bone ash, $p = 0.10$) (Fig., b). In a validation study examining the fluoride content between iliac crest bone and tumor-adjacent bone among cases, the Spearman correlation was 0.61 ($p < 0.001$).

There was no significant difference in the median fluoride concentration in bone between the matched osteosarcoma case and tumor control pairs ($N = 32$) (median = 804 vs. 714 mg F/kg of bone ash, $p = 0.63$) (Fig., c). When bone specimens from all cases ($N = 137$) and controls ($N = 51$) were included in an unmatched analysis, the median bone fluoride concentration in tumor-adjacent bone was significantly higher in controls than in cases (median = 754 vs. 611 mg F/kg of bone ash, $p = 0.01$) (Fig., d).

There were no differences in the results of the conditional and unconditional analyses; thus, the results of the unconditional analyses are reported given the increased power for detecting associations with bone fluoride. In the age- and sex-adjusted analysis, OR = 1.22 for an increase in bone fluoride from the 25th percentile (463.5 ppm) to the 75th percentile (943.3 ppm), representing an OR = 1.32 (95% CI, 0.58-3.03) for a 1-unit increase in the natural log of fluoride (ppm). After adjustment for age and gender, history of broken bones, other bone diseases, other cancer diagnoses, and history of receiving radiation prior to illness were significant covariates (see Appendix Table 2). The OR for log bone fluoride adjusted for these predictors, age, and gender was 1.23 (95% CI, 0.51-2.97) (Table 2). With a stepwise selection method to determine the final model adjusted for age and gender, a history of broken bones remained as a significant predictor, and the final adjusted OR was 1.33 (95% CI: 0.56-3.15) (Table 2).

In an analysis restricted to patients < 45 yrs old (123 cases; 30 controls), history of broken bones was the only predictor of osteosarcoma risk at the $p = 0.25$ level, adjusted for age and gender. However, the final risk-adjusted model included only age and gender (OR = 1.23, 95% CI: 0.48-3.16) (Table 2).

DISCUSSION

The results of the present study are similar to results of several other case-control studies that included histories of fluoride exposure based on community water fluoride concentrations or

from other fluoride sources, such as toothpaste and supplements (McGuire *et al.*, 1991; Gelberg *et al.*, 1995; Moss *et al.*, 1995).

The higher median fluoride concentration of controls compared with that of cases in this study is likely due to the fact that control patients tended to be older than the cases. In this study, fluoride content in bone in both cases and controls increased with age (moderate positive correlation, data not shown), which is similar to findings in studies that looked at the relationship between bone fluoride content and age (Parkins *et al.*, 1974; Eble *et al.*, 1992; Richards *et al.*, 1994).

Previous ecological and case-control studies that relied on historic residential information were limited in that they did not reflect the true exposure of fluoride at the individual level; thus, such studies are subject to the “ecological fallacy” (Aschengrau and Seage, 2003). They also did not take into account population mobility between fluoridated and non-fluoridated areas, or changes in population size and age (Freni and Gaylor, 1992), or potential confounders.

In this study, cases were all recruited from academic referral centers for bone cancer and thus were not a random sample of osteosarcoma patients. Controls were also bone cancer patients recruited from these same centers, and thus likely reflect the same source population as the cases. Although there was a difference in participation rates in the bone donation component, with 76% of the cases and 48% of the controls participating, it is unlikely that any enthusiasm for participation was related to bone fluoride levels.

Misclassification of exposure and/or outcome is always a concern in observational studies; however, given the laboratory measurement of fluoride exposure and the histologic confirmation of cases, misclassification bias is likely to be minimized in this study. The coefficient of variation for deer bone specimens, included in each batch of specimens as quality control, was 0.03, further decreasing the likelihood of substantial non-random misclassification bias.

There are also some potential drawbacks to the use of bone fluoride measurements. For example, if risk is related to exposures at a specific time in life, rather than total accumulated dose, this metric would not be optimal. Also, it is possible that fluoride concentrations in bone may be influenced by the disease, or that concentrations in tumor tissue are not representative of pre-disease levels. For this reason, we chose normal tumor adjacent bone from the surgical specimens instead of the tumor tissue. It is possible that bone metabolism in the vicinity of the tumor could also be disrupted, however, we chose other bone tumors as controls. If such a circumstance prevailed, it would likely affect the controls in a manner similar to the cases. To address this concern, we compared fluoride levels in bone adjacent to the tumor with those in bone from a distant site (iliac crest) from the same patients. While the absolute levels were different, as anticipated from the different kinds of bone involved, there was a highly significant rank order correlation between the fluoride levels from these 2 locations, lending some confidence to the validity of the adjacent bone measures (data not shown).

If fluoride levels were related to bone cancer in general, the current study design would be unable to detect this. There is no published evidence of such an association.

Table 1. Patient Characteristics of Osteosarcoma Cases (N = 137) and Tumor Controls (N = 51) Who Provided a Bone Specimen^a

		Cases (%) (N = 137) ^b	Tumor Controls (%) (N = 51) ^b	p Value ^c (Chi-square)
Site of enrollment	MGH, Boston, MA	24 (17.5)	7 (13.7)	0.58 ^d
	Creighton Univ./St. Joseph's, Omaha, NE	3 (2.2)	4 (7.8)	
	University of Nebraska, Omaha	12 (8.8)	1 (2.0)	
	University of Chicago	37 (27.0)	20 (39.2)	
	Rush Presbyterian, Chicago, IL	13 (9.5)	4 (7.8)	
	University of Florida, Gainesville	21 (15.3)	7 (13.7)	
	University of California, Los Angeles	15 (11.0)	7 (13.7)	
	Cleveland Clinic	10 (7.3)	1 (2.0)	
	Children's Nat'l Med. Ctr., Washington, DC	2 (1.5)	0 (0)	
Gender	Male	73 (53.3)	36 (70.6)	0.03
	Female	64 (46.7)	15 (29.4)	
Race/Ethnicity	White, Non-Hispanic	112 (81.8)	41 (80.4)	0.83 ^e
	Hispanic	12 (8.8)	1 (2.0)	
	Black, Non-Hispanic	8 (5.8)	6 (11.8)	
	Asian and Pacific Islander	3 (2.2)	0 (0)	
	Other	2 (1.5)	3 (5.9)	
Patient's education	Less than high school	74 (54.0)	18 (35.3)	0.07
	HS/equivalent/post-HS training/ some college	43 (31.4)	22 (43.1)	
	College or post-grad	20 (14.6)	11 (21.6)	
	Missing	0 (0)	0 (0)	
Mother's education	Less than high school	14 (10.2)	9 (17.7)	0.28 ^f
	HS/equivalent/post-HS training/ some college	85 (62.0)	30 (58.8)	
	College or post-grad	36 (26.3)	10 (19.6)	
	Missing	2 (1.5)	2 (3.9)	
Father's education	Less than high school	15 (11.0)	13 (25.5)	0.02
	HS/equivalent/post-HS training/ some college	72 (52.6)	17 (33.3)	
	College or post-grad	44 (32.1)	16 (31.4)	
	Missing	6 (4.4)	5 (9.8)	
Combined household income	≤ \$40,000	54 (37.2)	18 (35.3)	0.84
	\$40,001 - \$60,000	27 (19.7)	13 (25.5)	
	> \$60,000	38 (27.7)	13 (25.5)	
	Missing	18 (13.1)	7 (13.7)	
Urban	Ever lived in urban area	119 (86.9)	47 (92.2)	0.32
	Never lived in urban area	18 (13.1)	4 (7.8)	
Age (yrs)	0 – 14	37 (27.0)	9 (17.7)	< 0.001
	15 – 29	72 (52.6)	12 (23.5)	
	30 – 44	13 (9.5)	9 (17.7)	
	45 and older	15 (10.9)	21 (41.2)	
Median age (yrs)	Overall	17.6	41.3	< 0.001 ^g
	Males	17.0	42.0	< 0.001 ^g
	Females	17.0	39.0	0.02 ^g

^aThere were 194 patients who provided tumor-adjacent bone specimens; however, fluoride concentrations from specimens (five cases and one tumor control) were below < 100 mg F/kg bone ash and were not included in the analysis. In total, 188 patients were used in the analysis.

^bPercentages do not add up to 100 because of rounding.

^cChi-square testing differences between cases and tumor controls with bone specimens.

^dFor the comparisons, patients from MGH and Children's Nat'l Med. Ctr. were grouped together, since they are in the same region; patients from Creighton Univ. and the Univ. of Nebraska were grouped together, since they are in the same city; and patients from the Univ. of Chicago and Rush Presbyterian were grouped together, since they are in the same city.

^eComparing White, Non-Hispanics with all other racial/ethnic groups.

^fFisher's exact test.

^gWilcoxon rank sum.

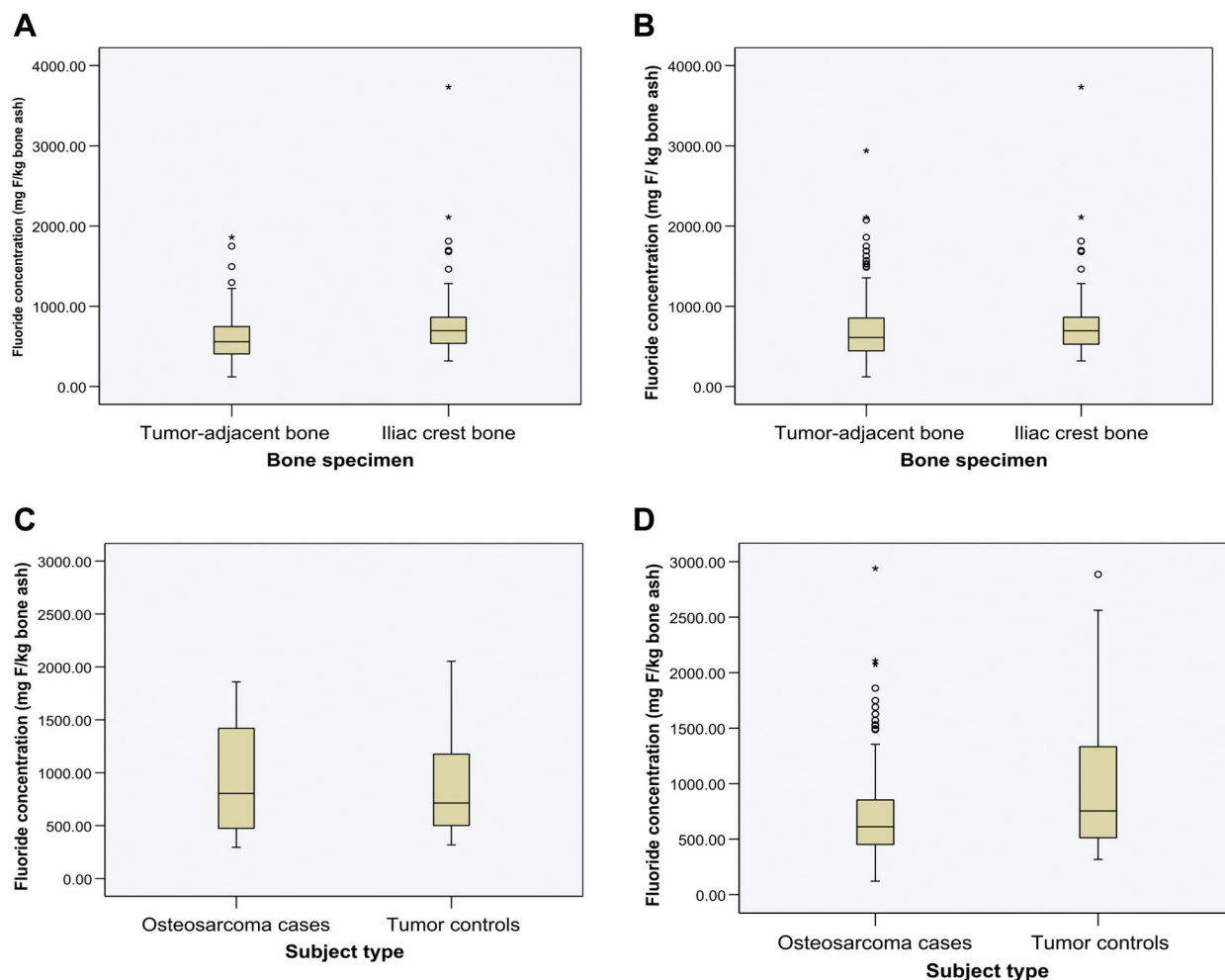


Figure. Box plots of interquartile range (IQR), range, and median fluoride concentrations (mg/kg bone ash). **(A)** Tumor-adjacent bone vs. iliac crest bone among cases that contributed both samples: IQR = 407.0-746.5 vs. 538.0-863.0; range = 121.0-1859.5 vs. 318.5-3732.5; median = 558.0 vs. 696.5 ($p < 0.001$, Wilcoxon signed-rank test). **(B)** Tumor-adjacent bone vs. iliac crest bone among all cases: IQR = 444.5-853.0 vs. 527.0-863.5; range = 121.0-2939.0 vs. 318.5-3732.5; median = 611.0 vs. 696.5 ($p = 0.10$, Wilcoxon rank-sum test). **(C)** Tumor-adjacent bone in matched cases vs. tumor controls: IQR = 474.8-1419.5 vs. 501.0-1176.5; range = 296.0-1859.5 vs. 317.5-2053.5; median = 803.8 vs. 714.3 ($p = 0.63$, Wilcoxon signed-rank test). **(D)** Tumor-adjacent bone among all cases vs. tumor controls: IQR = 451.5-853.0 vs. 501.5-1359.0; range = 121.0-2939.0 vs. 317.5-2885.0; median = 611.0 vs. 754.0 ($p = 0.0008$, Wilcoxon rank-sum test).

The major advantage of this study is the use of bone fluoride concentrations as the measure of fluoride exposure, rather than estimating fluoride exposure in drinking water (Bassin *et al.*, 2006). Since 99% of the body burden of fluoride is located in calcified tissues, and fluoride concentration is dependent upon the amount and duration of exposure as well as the rate of bone turnover (Turner *et al.*, 1993), if chronic fluoride intake was a risk factor for osteosarcoma, then it would be reasonable to expect that cases would have significantly higher bone fluoride concentrations than tumor controls. This study did not demonstrate an association between fluoride levels in bone and osteosarcoma.

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CWD has written reviews of the literature for several companies that sell, reimburse for, or do research on preventive dentistry products, most notably GlaxoSmithKline, Colgate-Palmolive, Dentsply, Quintile, Delta Dental Plans, and the United States Public Health Service (USPHS). CH has done limited consulting with Procter & Gamble. All other authors have no conflict of interest to disclose.

Table 2. Odds Ratios and 95% CI for a 1-unit Increase in Natural Log of Fluoride Concentration (ppm) in Bone and Risk of Osteosarcoma: (A) for All Osteosarcoma Cases and Tumor Controls and (B) among Patients under 45 Years of Age

	OR	95% CI	p Value
A. For all osteosarcoma cases and tumor controls			
Age- and sex-adjusted model ^a	1.32	(0.58, 3.03)	0.51
Fully adjusted model ^b	1.23	(0.51, 2.97)	0.65
Risk-adjusted model ^c	1.33	(0.56, 3.15)	0.58
B. Among patients younger than 45 yrs old			
Age- and sex-adjusted model ^a	1.23	(0.48, 3.16)	0.67
Fully adjusted model ^d	1.27	(0.49, 3.35)	0.62
Risk-adjusted model ^e	1.23	(0.48, 3.16)	0.67

^aIncludes only age and gender.

^bIncludes all variables that were significant at $p \leq 0.25$ in the age- and sex-adjusted analysis (history of broken bones, other cancers, other bone diagnosis, and received radiation prior to illness), plus age and gender.

^cIncludes history of broken bones, plus age and gender.

^dIncludes history of broken bones (significant at $p \leq 0.25$ in the age- and sex-adjusted analysis), plus age and gender.

^eIncludes age and gender and no other variables.

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RESEARCH REPORTS

Clinical

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APPENDIX

METHODS

Study Population

Incident cases and matched controls were selected from 10 hospitals across the US between 1993 and 2000. Patients were identified by physicians in the orthopedic departments of participating institutions and had to be either living in the US or US citizens temporarily living outside the US during the enrollment period to be considered eligible.

The original plan was to match participants based on age, gender, and geographic ring of last residence. The matching on geographic ring was planned mainly to match SES-related factors, and such matching could occasionally result in matching on fluoride (not always, since most people lived in several residences over their lifetime, and there could be different fluoride levels within one geographic ring). Not all cases and controls

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provided bone specimens, so the number of matched pairs for which both the case and control provided bone was very limited (32 of 142 cases, 23%). Importantly, due to the logistical difficulty in finding matched tumor controls, the matching criteria were relaxed during the recruitment, and only 14 of the 32 case-control pairs used in this manuscript (44%) were actually matched on geographic ring.

Tumor and orthopedic controls were matched to cases on gender, age within ± 5 yrs, and initially by residence as defined with a geographic area/ring defined by distance from their residence at the time of enrollment to the respective medical center. The postal ZIP Code prefix (the first 3 digits) for each residence was used to create geographic rings. For example, for the Boston, MA, area, "Ring 1" included ZIP Code prefixes from 014xx through 027xx, which included all of eastern and central MA. The geographic criterion was relaxed later in the enrollment process, because the catchment area for patient referrals with musculoskeletal tumors was broader than that for patients with benign orthopedic conditions and trauma.

Appendix Table 1. Patient Characteristics of Patients Who Had Bone Sample (N = 188) and Patients Who Did Not Have Bone Sample (N = 102)

		Patients Who Had Bone Sample	Patients Who Did Not Have Bone Sample	p Value (Chi-square)
		(N = 188) ^a	(N = 102) ^a	
Site of enrollment	MGH, Boston, MA	31 (16.5)	25 (24.5)	< 0.001 ^b
	Creighton Univ./St. Joseph's, Omaha, NE	7 (3.7)	0 (0.0)	
	University of Nebraska, Omaha	13 (6.9)	1 (1.0)	
	University of Chicago	57 (30.3)	4 (3.9)	
	Rush Presbyterian, Chicago, IL	17 (9.0)	16 (15.7)	
	University of Florida, Gainesville	28 (14.9)	49 (48.0)	
	University of California, Los Angeles	22 (11.7)	0 (0.0)	
	Cleveland Clinic	11 (5.9)	3 (2.9)	
	Children's Nat'l Med. Ctr., Washington, DC	2 (1.1)	4 (3.9)	
Gender	Male	109 (58.0)	58 (56.9)	0.85
	Female	79 (42.0)	44 (43.1)	
Race/Ethnicity	White, Non-Hispanic	153 (81.4)	86 (84.3)	0.53 ^c
	Hispanic	13 (6.9)	4 (3.9)	
	Black, Non-Hispanic	14 (7.4)	9 (8.8)	
	Asian and Pacific Islander	3 (1.6)	2 (2.0)	
	Other	5 (2.7)	1 (1.0)	
Patient's education	Less than high school	92 (48.9)	41 (40.2)	0.21
	HS/equivalent/post-HS training/some college	65 (34.6)	46 (45.1)	
	College or post-grad	31 (16.5)	15 (14.7)	
Mother's education	Less than high school	23 (12.2)	30 (29.4)	< 0.001 ^d
	HS/equivalent/post-HS training/some college	115 (61.2)	61 (59.8)	
	College or post-grad	46 (24.5)	11 (10.8)	
	Missing	4 (2.1)	0 (0.0)	
Father's education	Less than high school	28 (14.9)	24 (23.5)	0.04
	HS/equivalent/post-HS training/some college	89 (47.3)	55 (53.9)	
	College or post-grad	60 (31.9)	20 (19.6)	
	Missing	11 (5.9)	3 (2.9)	
Combined household income	≤ \$40,000	72 (38.3)	44 (43.1)	0.67
	\$40,001 - \$60,000	40 (21.3)	17 (16.7)	
	> \$60,000	51 (27.1)	25 (24.5)	
	Missing	25 (13.3)	16 (15.7)	
Urban	Ever lived in urban area	166 (88.3)	95 (93.1)	0.19
	Never lived in urban area	22 (11.7)	7 (6.9)	
Age (yrs)	0 – 14	46 (24.5)	18 (17.7)	0.13
	15 – 29	84 (44.7)	39 (38.2)	
	30 – 44	23 (12.2)	20 (19.6)	
	45 and older	35 (18.6)	25 (24.5)	
Median age (yrs)	Overall	19.0	25.0	0.02 ^e
	Males	20.0	25.0	0.29 ^e
	Females	18.0	25.0	0.02 ^e

^a Percentages do not add up to 100 because of rounding.

^b For the comparisons, patients from MGH and Children's Nat'l Med. Ctr. were grouped together, since they are in the same region; patients from Creighton Univ. and the Univ. of Nebraska were grouped together, since they are in the same city; and patients from the Univ. of Chicago and Rush Presbyterian were grouped together, since they are in the same city.

^c Comparing White, Non-Hispanics with all other racial/ethnic groups; ^d Fisher's exact test; ^e Wilcoxon rank-sum test.

No statistically significant differences were found between the patients who provided bone specimens and patients who did not with respect to gender, race/ethnicity, patient's education, combined household income, and urban area. There were differences in parental level of education, with mothers and fathers of patients with bone specimens tending to be more highly educated ($p < 0.001$ and $p = 0.04$, respectively), and between enrollment site, where a greater proportion of patients from UCLA, Cleveland, and the Omaha and Chicago sites provided bone specimens ($p < 0.001$). Patients with bone samples were overall slightly younger compared with those without samples ($p = 0.02$). Female cases and controls who provided bone were younger, on average, than females who did not ($p = 0.03$); however, no difference in ages of male cases and controls was observed ($p = 0.29$).

Appendix Table 2. Odds Ratios and 95% Confidence Intervals for Potential Confounders with an Age- and Sex-adjusted Model

	N ^a	OR	95% CI	p Value
Age (yrs)	188	0.95	(0.93, 0.97)	< 0.001
Gender	188	0.48	(0.24, 0.95)	0.03
Covariates below are adjusted for age and gender				
Ever had a broken bone ^b	188	0.43	(0.20, 0.90)	0.03
Other cancers ^b	187	4.59	(0.83, 25.35)	0.08
Other bone diagnoses ^b	187	2.92	(0.73, 11.74)	0.13
Ever received radiation prior to illness ^b	187	0.56	(0.25, 1.24)	0.15
Father's education	188			0.44
Less than high school		1.00	-	-
HS/equivalent/post-HS training/some college		1.71	(0.59, 4.90)	0.32
College or post-grad		0.89	(0.29, 2.73)	0.84
Missing		0.96	(0.18, 4.97)	0.96
Site of enrollment	188			0.60
MGH and Children's Nat'l Med. Ctr.		1.00	-	-
Creighton Univ. and Univ. of Nebraska, Omaha		0.61	(0.15, 2.50)	0.49
Univ. of Chicago and Rush Presbyterian		0.61	(0.20, 1.85)	0.38
Univ. of Florida		1.28	(1.33, 4.97)	0.72
UCLA		0.63	(0.16, 2.48)	0.51
Cleveland Clinic		2.68	(0.27, 27.16)	0.40
Distance (<i>per</i> 100 miles)		0.97	(0.87, 1.09)	0.59
Combined household income	188			0.63
≤ \$40,000		1.00	-	-
\$40,000 - \$60,000		0.61	(0.24, 1.59)	0.31
> \$60,000		1.04	(0.42, 2.59)	0.94
Missing		0.61	(0.19, 2.00)	0.42
Patient's education	188			0.49
Less than high school		Ref		
HS/equivalent/post-HS training/some college		1.40	(0.54, 3.65)	0.49
College or post-grad		2.15	(0.62, 7.47)	0.23
White, Non-Hispanic	188	1.17	(0.47, 2.90)	0.73
Living in an urban area at time of enrollment	187	1.02	(0.30, 3.41)	0.98
Mother's education	188			0.95
Less than high school		Ref		
HS/equivalent/post-HS training/some college		0.79	(0.25, 2.44)	0.68
College or post-grad		0.68	(0.18, 2.59)	0.57
Missing		0.62	(0.05, 7.61)	0.71

^a In some cases, the total number of patients does not equal 188, due to patients who were missing data for a given variable.

^b Variables significant at the $p \leq 0.25$ level in the age- and sex-adjusted analysis.